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General Bayesian approach to design clinical trials with time-to-event outcomes

James McGree Professor of Statistics School of Mathematical Sciences Queensland University of Technology

james.mcgree@qut.edu.au | www.jamesmcgree.com | @j_mcgree

Collaborators: Overstall, Mahar, Jones, Snelling

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James McGree AusTRIM 2023

Outline of talk

Introduction

- Motivation: Re-design of ORVAC
- Background:
 - (Standard) Bayesian inference
 - General Bayesian methods for inference
 - (Standard) Bayesian design for clinical trials
- Extension: General Bayesian design for clinical trials
- Trial simulation for re-designing ORVAC
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- COVID-19 brought an unprecedented push to undertake clinical trial assessments as quickly as possible
- Adaptive trials are particularly appealing in this regard
- Based on information from accruing data, e.g. trial can stop early due to success or futility, treatment can be dropped if ineffective or unsafe, etc
- Compared to non-adaptive trials, can complete sooner, cost less to run and reduce the number of participants on inferior treatments
- Potential benefits for participants and trialists

- Typically adaptive trials are designed via trial simulation
- Assumptions about the data generating process needed
- e.g. Size of treatment effect and distribution of data
- Misspecification can lead to a poorly designed trial e.g. insufficient power
- Could have consequences in terms of conclusions drawn from trial
- Propose to address this, at least partly, through general Bayesian methods

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- ORVAC: Optimising Rotavirus Vaccine in Aboriginal Children
- Double-blind, randomised, placebo-controlled Bayesian adaptive clinical trial
- Assess effectiveness of a third dose of Rotarix rotavirus vaccine in Australian Indigenous infants in providing improved protection (versus usual care) against gastroenteritis
- The third dose was active treatment, usual care was matched placebo.
- Randomly allocated to participants at a ratio of 1:1

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- There were two primary outcomes:
 - Anti-rotavirus IgA seroconversion, defined as serum anti-rotavirus IgA ≥ 20 U/ml 28 to 55 days post Rotarix/placebo
 - Time from randomisation to presentation of acute gastroenteritis or acute diarrhoea illness (t2e outcome)
- Eligible for enrolment: Aged between 6 and 12 months
- Followed up to 36 months of age
- The maximum sample size: 1,000 participants
- Planned analyses occurred after 250 participants then every 50 thereafter.
- At each planned analysis, pre-specified decision rules: effectiveness and futility.

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- To design trial, simulation was used
- Assumed proportional hazards where baseline hazard from exponential or Weibull
- Hazard function of the form:

$$h_i(t) = \psi_i h_0(t),$$

where t is time since enrolment, $\psi_i = \exp(x_i\beta)$, x_i is treatment, β is treatment effect and $h_0(t)$ is the baseline hazard function for participant *i*.

 Baseline hazard: Assumed constant under exp, and monotonically increasing or decreasing under Weibull

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• All inference based on the posterior distribution of θ i.e.

 $p(\theta|\mathbf{y}, \mathbf{x}) \propto p(\theta)p(\mathbf{y}|\theta, \mathbf{x}),$

where $p(\theta)$ is the prior, $p(\mathbf{y}|\theta, \mathbf{x})$ is the likelihood, $\mathbf{y} = (y_1, \dots, y_N)^t$ and $y_i = (t_i, c_i)$.

- Make probability statements about treatment effect
- For t2e outcomes, likelihood can be formed based on hazard function
- E.g. Specify baseline hazard and assume proportional hazards between treatments

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- Specifically, $f(t|\theta) = h(t)S(t|\theta)$, where $S(t|\theta) = \exp(-H(t))$ and $H(t) = \int_0^t h(t)dt$
- To construct likelihood for ORVAC, participant i yield a t2e outcome from treatment allocation x_i.
- Assuming each observation is conditionally independent:

$$p(\mathbf{y}|\boldsymbol{\theta}, \mathbf{x}) = \prod_{i=1}^{N} f(t_i|\boldsymbol{\theta}, x_i)^{c_i} S(t_i|\boldsymbol{\theta}, x_i)^{1-c_i},$$

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where c_i denotes censored or not for *i*th participant (uninformative right censoring).

General Bayesian inference

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Inference based on a general loss function i.e.

$$p(\theta|\mathbf{y}, \mathbf{x}) \propto p(\theta) \exp[-wl(\theta, \mathbf{y}, \mathbf{x})],$$

Trial simulation

where $w \ge 0$ and $l(\theta, \mathbf{y}, \mathbf{x})$ is the loss function.

Extension

- Standard Bayesian inference is a special case
- For continuous and distinct outcomes $t_{(1)}, \ldots, t_{(J)}$, negative partial log-likelihood as the loss function (with w = 1) i.e.

$$l(\boldsymbol{ heta}, \boldsymbol{y}, \boldsymbol{x}) = -\sum_{j=1}^{J} \log rac{\exp(x_{(j)} \boldsymbol{ heta})}{\sum_{r \in R_j} \exp(x_r \boldsymbol{ heta})},$$

 R_j is risk set (set of participants not yet responded and not censored) at the *j*th ordered event time $t_{(j)}$ for a participant defined by $x_{(j)}$

Note: Baseline hazard nuisance/undefined



References

Bayesian adaptive trials

- Typically trial design assessed via trial simulation
- Requires pre-specifying decision rules and what adjustments can be made
- When interim analyses will occur
- Used to provide insight into performance e.g. empirical estimates of power and type-1 error
- Adjust design to optimise summarise in some way

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Bayesian adaptive trials

Simulate adaptive trial

- 1: Initialise $p(\theta)$, h(t), treatment effect
- 2: for k = 1 : K do
- 3: Enrol participants
- 4: Randomly assign treatments **x**
- 5: Simulate observations based on x and h(t)
- 6: if planned interim analysis then
- 7: Evaluate decision rules
- 8: **if** any decision rule met **then**
- 9: Stop or adjust trial
- 10: end if
- 11: end if
- 12: end for
- 13: Summarise simulated trial

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ORVAC - Effectiveness decision rule

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Expectation of declaring trial effectiveness based on:

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 $P(eta < 0 | oldsymbol{y}, oldsymbol{x}) > 0.97$

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Expectation taken over data from participants who have not responded and are not censored i.e.

$$\left[\int_{\mathcal{Z}} \mathcal{I}(P(\beta < 0 | \boldsymbol{y}, \boldsymbol{z}, \boldsymbol{x}) > 0.97) p(\boldsymbol{z} | \boldsymbol{y}, \boldsymbol{x}) \mathrm{d}\boldsymbol{z}\right] > 0.9$$

No analytic solution available but can use Monte Carlo:

$$pprox 1/B\sum_{b=1}^{B}\mathcal{I}(P(eta < 0|m{y},m{z}_b,m{x}) > 0.97)$$

where z_b is simulated from posterior predictive distribution.

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ORVAC - Effectiveness decision rule

Evaluate effectiveness decision rule

- 1: Initialise $\boldsymbol{y}, \boldsymbol{x}, p(\boldsymbol{\theta}|\boldsymbol{y}, \boldsymbol{x})$
- 2: for b = 1 : B do
- 3: Simulate posterior predictive data z_b for enrolled participants where outcome not observed and not censored
- 4: Update posterior distribution
- 5: Evaluate trial effectiveness rule

$$oldsymbol{\Lambda}^b = \mathcal{I}(P(eta < 0 | oldsymbol{y}, oldsymbol{z}_b, oldsymbol{x}) > 0.97)$$

- 6: end for
- 7: $\delta_e = \frac{1}{B} \sum_{b=1}^{B} \Lambda^b$
- 8: if $\delta_e > 0.90$ then
- 9: Stop trial
- 10: end if

ORVAC - Futility decision rule

Background

- Expectation of declaring trial effectiveness
- Expectation taken over:

Motivation

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- Data from participants who have not responded and are not censored
- Data from participants not yet enrolled, i.e.

Extension

$$\left[\int_{\mathcal{Z}}\sum_{\boldsymbol{v}\in\mathcal{V}}\mathcal{I}(P(\beta<0|\boldsymbol{y},\boldsymbol{z},\boldsymbol{x},\boldsymbol{v})>0.97)p(\boldsymbol{z}|\boldsymbol{y},\boldsymbol{x},\boldsymbol{v})p(\boldsymbol{v})\mathrm{d}\boldsymbol{z}\right]<0.05$$

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No analytic solution available but can use Monte Carlo:

$$pprox 1/B\sum_{b=1}^{B}\mathcal{I}(P(eta < 0|m{y},m{z}_b,m{x},m{v}_b) > 0.97)$$

where z_b is simulated from posterior predictive distribution, and v_b from randomisation scheme.



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ORVAC - Futility decision rule

Evaluate futility decision rule

- 1: Initialise $\boldsymbol{y}, \boldsymbol{x}, p(\boldsymbol{\theta}|\boldsymbol{y}, \boldsymbol{x})$
- 2: for b = 1 : B do
- 3: Simulate treatment allocation v_b
- 4: Simulate posterior predictive data z_b for enrolled participants and for participants who are yet to enrol
- 5: Update posterior distribution
- 6: Evaluate trial effectiveness rule

$$oldsymbol{\Lambda}^b = \mathcal{I}(oldsymbol{P}(eta < 0 | oldsymbol{y}, oldsymbol{z}_b, oldsymbol{x}, oldsymbol{v}_b) > 0.97)$$

- 7: end for
- 8: $\delta_f = \frac{1}{B} \sum_{b=1}^{B} \Lambda^b$
- 9: if $\delta_f < 0.05$ then
- 10: Stop trial
- 11: end if

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General Bayesian adaptive trials

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Trial simulation: Need to make assumptions about the data generating process

Trial simulation

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We seek robustness to the misspecification of these assumptions

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- Our approach: Propose all inference conducted within a general Bayesian framework.
- Evaluate trial designs based on the partial likelihood i.e.
 Baseline hazard remains nuisance/undefined
- However, loss function need not be linked to the data generating process, so how to simulate data?
- i.e. Cannot generate data from partial likelihood (only)

General Bayesian adaptive trials

- To simulate data, propose to consider a *super model*
- Formulated such that it can describe a wide range of data sets

References

- Here, a wide variety of hazard functions such as constant, monotonic and non-monotonic function
- Could also extend to consider more complex censoring mechanisms
- Super model is out of the scope of inference
- So not desirable to estimate a treatment effect based on this model
- e.g. most likely overparameterised

General Bayesian adaptive trials

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Here, define super model based on a cubic spline representation of the baseline hazard function i.e.

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$$h_0(t) = \sum_{q=1}^Q \xi_q g_q(t),$$

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where ξ_q are parameters and $g_q(t) = t^q$ are the basis functions, for q = 1, ..., Q.

- Very flexible: Constant, monotonic and non-monotonic forms
- Use this model for simulation, and use partial likelihood for inference

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- Use trial simulation to explore new approach to re-design for ORVAC
- Models for data generation in trial simulation:
 - Exponential PHs model
 - Weibull PHs model
 - Super model
- Models for estimating treatment effect in trial simulation:
 - Exponential PHs model
 - Weibull PHs model
 - General Bayesian model
- Nine combinations
- Vaguely informative priors used throughout

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- Initially, different values for treatment effect i.e. $\beta \in \{0, -0.075, -0.125, -0.175, -0.25, -0.5\}$
- Consider specific forms for baseline hazard

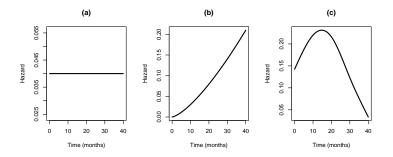


Figure: Assumed hazard functions for the (a) Exponential, (b) Weibull and (c) Super model.



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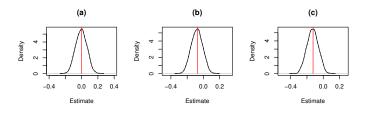
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Estimation results



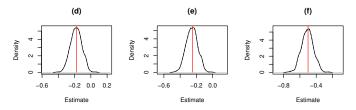


Figure: Data from exponential model, then fit Exponential (\cdots) , general Bayesian (-) and Weibull (--).

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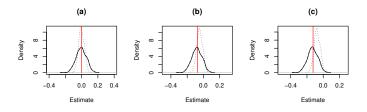
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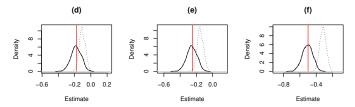


Figure: Data from Weibull model, then fit Exponential (\cdots) , general Bayesian (-) and Weibull (--).

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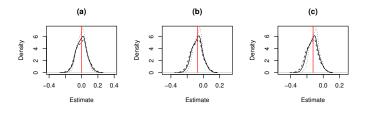
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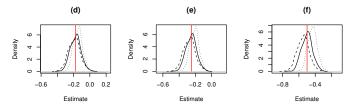
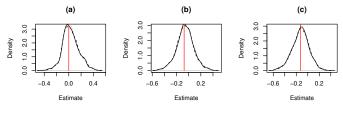


Figure: Data from super model, then fit Exponential (\cdots) , general Bayesian (-) and Weibull (--).



Trial simulation results



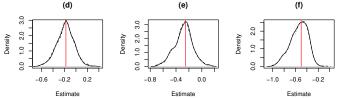
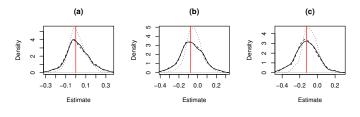


Figure: Data from exponential model, then fit exponential (\cdots) , general QUT Bayesian (-) and Weibull (-).

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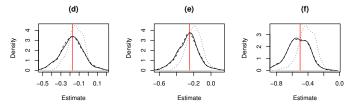
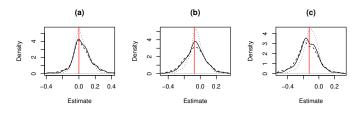


Figure: Data from Weibull model, then fit exponential (\cdots) , general Bayesian (-) and Weibull (-). → < ∃→ iscussion 00 References 0

Trial simulation results



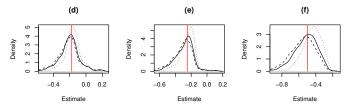


Figure: Data from super model, then fit exponential (\cdots) , general Bayesian (-) and Weibull (--).

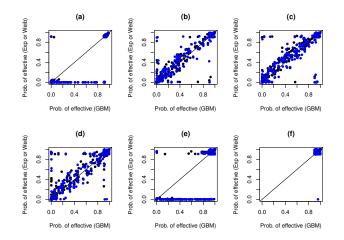


Figure: Data from exponential model, then fit general Bayesian, exponential (black) and Weibull (blue) models.

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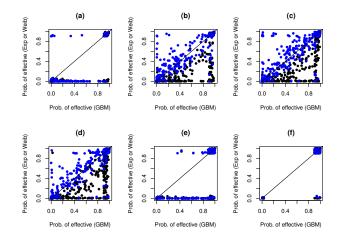


Figure: Data from Weibull model, then fit general Bayesian, exponential (black) and Weibull (blue) models.

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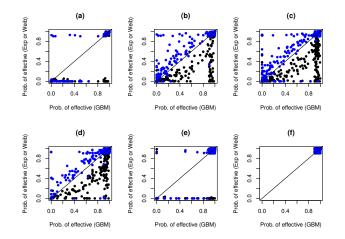


Figure: Data from super model, then fit general Bayesian, exponential (black) and Weibull (blue) models.

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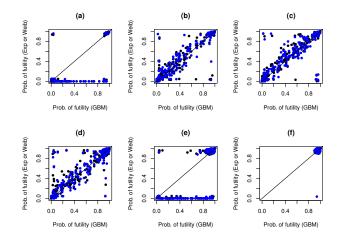


Figure: Data from exponential model, then fit general Bayesian, exponential (black) and Weibull (blue) models.

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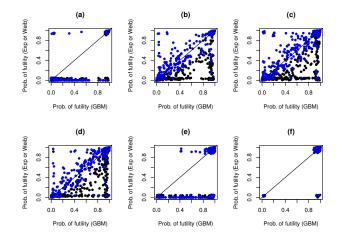


Figure: Data from Weibull model, then fit general Bayesian, exponential (black) and Weibull (blue) models.

Trial simulation

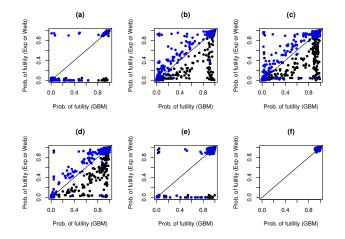


Figure: Data from super model, then fit general Bayesian, exponential (black) and Weibull (blue) models.



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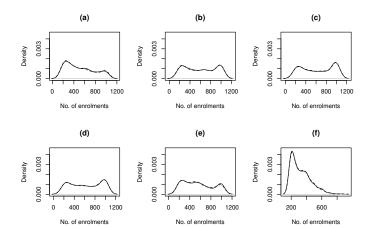
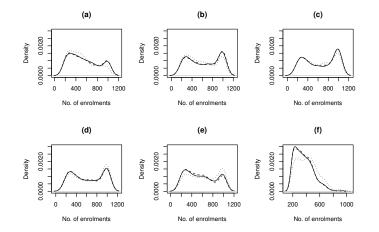


Figure: Data from exponential model, then fit exponential (\cdots) , general QUT Bayesian (-) and Weibull (-) models. ⇒ →



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Figure: Data from Weibull model, then fit exponential (\cdots) , general Bayesian (-) and Weibull (--) models.



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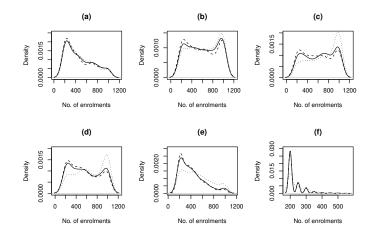


Figure: Data from super model, then fit exponential (\cdots) , general Bayesian (-) and Weibull (--) models.



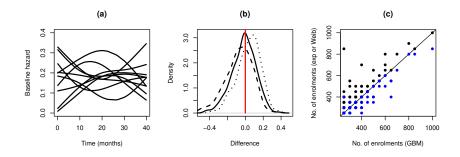
Extended adaptive trial results

- Next, explore the performance of models under a range of baseline hazard functions from the super model
- Form a prior for the super model
- i.e. Fix number of knots and knot position, then random generate corresponding values from U[0, 0.4] independently
- Prior on treatment effect: $\beta \sim U(-0.75, -0.25)$
- Assess performance as before

Extended adaptive trial results

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Figure: (a) 10 realisations of baseline hazard from super model; and (b) distribution of the difference between the true and estimated treatment effect under the exponential (\cdots) , general Bayesian (-) and Weibull (--) models, and (c) number of enrolments under the general Bayesian model compared to the exponential (black) and Weibull (blue) models.



References

- Exp and Weibull provide slightly biased estimates of treatment
- Median difference of 0.0503 and −0.0607, respectively.
- General Bayesian model relatively unbiased estimate of treatment with mean difference of −0.0087
- Exp/Weibull, on average, a larger/fewer number of enrolments compared to general Bayesian model
- Mean difference of 18/-5.2 enrolments.
- No appreciable differences in probabilities of success and futility

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- Proposed a general Bayesian method to design adaptive trials
- For time-to-event outcomes, not required to specify the specific data generating model
- Led to trial design that was robust to baseline hazard function
- Appears useful e.g. assuming the wrong baseline hazard function can lead to over/underestimation, and shorter/longer trials
- Seems preferable to base designs on flexible models



- Scope to extend to other outcomes e.g. Overdispersed counts, remove influence of outliers, etc
- More broadly, seek approaches to reduce reliance on assumptions e.g. PH.
- Need computationally efficient approaches to determine w
- Robust approach in other settings? e.g. GAMs, Gaussian Process?



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