Principles of experimental design for big data analysis

James McGree Associate Professor of Statistics School of Mathematical Sciences Queensland University of Technology

james.mcgree@qut.edu.au | www.jamesmcgree.com @j_mcgree

Joint work with Drovandi, Holmes, Mengersen, Richardson and Ryan

ດມາ

• □ ▶ • • □ ▶ • □ ▶ • • □ ▶

Outline of talk

- Introduction/motivation
- Overview of Bayesian experimental design
- Experimental design in the context of big data
- Our proposed approach/algorithm
- Motivating examples:
 - 1 Mortgage default
 - 2 Accelerometer data
- Conclusions and future work

QUI

イロト イヨト イヨト

- Massive volumes of data being collected at an accelerating pace
- Traditional measurements are now complemented with digital data obtain from, e.g., images, text, audio, sensors, etc
- Such data have the potential to inform important problems in health, science, business, engineering, however....
- Size, complexity and quality makes these data sets difficult to process and analyse

ດມາ

- Often computationally prohibitive to store and manage such data sets on a single computer
- Generally require high performance computing
- Similarly, the analyses of these data requires computational and statistical techniques that exceeds typical capacity
- New technological and/or methodological methods are needed
- This motivates the development of new statistical methods for inference that accounts for the characteristics of the data, adjusts for potential bias/data gaps, and appropriately handles storage and analysis needs.

QD

Some methods to address challenges of managing, modelling and analysing big data

- Divide-and-conquer or divide-and-recombine methods (Xi et al., 2010 and Guhaa et al., 2012)
- Similar methods have proposed such as consensus Monte Carlo (Scott, Blocker and Bonassi, 2013) and bag of little bootstraps (Kleiner et al., 2014)
- Others have studied the properties of MCMC subsampling algorithms (Bardenet, Doucet and Holmes, 2014, 2015)
- Approaches for optimal subsampling (Wang et al., 2017/18)
- R packages for big data analyses (Wang et al., 2015)
- And many other developments....

- Recent(-ish) reviews are given by Fan, Han and Lui (2014) and Wang et al. (2015)
- Despite many advantages of the above (and other approaches), there is agreement that challenges remain
- For example
 - Reduce spurious correlations/patterns
 - Continuing to improve computational and algorithmic efficiency and stability
 - Accommodating heterogeneity and statistical biases associated with combining data from different sources using different technologies
- Given the acceleration of size and diversity of big data, potentially these will remain as stumbling blocks for the foreseeable future.



Proposed approach

- Propose an approach that may overcome some of these issues
- Targeted towards regression models with large N observations and small to moderate sized p predictors
- Depending on the analysis aim, use methods from optimal experimental design to make the inference problem more computationally tractable
- Instead of analysing the whole data set, a retrospective sample may provide enough information to address the analysis aim
- The specific sample is drawn in accordance with an experimental design based on a specific analysis question

QUT

・ロト ・ 聞 ト ・ ヨ ト ・ ヨ ト

The analysis is then based on this subsample

Proposed approach

- Thus, the big data challenge is being able to extract the design from the data set
- This is much more tractable
- Modelling problem reduces to a (near) designed analysis
- Should result in less correlation between predictors, potentially less spurious correlations and patterns, etc

ດມາ

Proposed approach

- Many big data inferential goals for which our approach is applicable
- Goals for which design principles and corresponding utility functions are well established
- Examples include estimation, testing significance of parameter values, prediction, identification of relationships, variable selection, etc
- Potential to consider such an approach for use within divide-and-conquer type algorithms and/or in sequential learning
- Potential to use this approach to evaluate the quality of the data including potential biases and data gaps



Background - Bayesian inference

• When interested in estimating θ

 $p(\theta|y,d) \propto p(\theta) p(y|\theta,d),$

where $p(\theta)$ is prior and $p(y|\theta, d)$ is the likelihood.

■ When interested in model choice, suppose *K* models are being considered, with *m* = 1,...,*K*

$$p(\theta_m|y, m, d) = rac{p(\theta_m|m)p(y|\theta_m, m, d)}{Z_m},$$

where $Z_m = \int_{\theta_m} p(\theta_m | m) p(y | \theta_m, m, d) d\theta_m$.

 $Z_m \propto p(m|y, d)$, so pick the model with the largest Z_m .

イロト 不得 とくほとく ヨト

Background

Two main challenges:

- 1 Approximating the expected utility;
- 2 Maximising the utility.
- Maximise expected utility $d^* = \arg \max_d u(d)$, where

$$u(d) = \sum_{m=1}^{K} p(m) \int_{y} u(d, y, m) p(y|d, m) dy.$$

- u(d, y, m) is some measure of information gained from d given model m and observed data y.
- Importantly, u(d, y, m) is typically a function of $p(\theta_m | y, m, d)$.

QUI

Background

- u(d) typically cannot be solved analytically
- Can be approximated using Monte Carlo integration

$$u(d) \approx \sum_{m=1}^{K} p(m) \frac{1}{B} \sum_{b=1}^{B} u(d, y_{mb}, m),$$

where $y_{mb} \sim p(y|\theta_{mb}, m, d)$ and $\theta_{mb} \sim p(\theta|m)$.

- Typical Bayesian utility is the KLD between the prior and the posterior.
- Hence, B posterior distributions need to be approximated or sampled from to approximated u(d).

ດມາ

イロト 不聞 とくほ とくほう

Computationally challenging task.

Entropy

- Suppose discrete random variables X and Y
- Define entropy:

$$H(X) = -\sum_{x \in X} f(x) \log f(x)$$

Define conditional entropy:

$$H(X|Y) = -\sum_{y \in \mathcal{Y}} f(y) \sum_{x \in \mathcal{X}} f(x|y) \log f(x|y)$$

Mutual information between X and Y is defined as:

$$I(X;Y) = H(X) - H(X|Y)$$

Design: Focus on expected change in entropy of X upon observing data Y.



Utility functions derived from mutual information

Estimation (KLD, Shannon, 1948)

$$u_P(d, y, m) = \int_{\theta} p(\theta|m, y, d) \log p(y|\theta, m, d) d\theta - \log p(y|m, d)$$

 Model discrimination (Box and Hill, 1967; Drovandi, McGree, Pettitt, 2014)

$$u_M(d, y, m) = \log p(m|y, d)$$

- Difficult to approximate $\log p(y|m, d)$ (and $\log p(m|y, d)$)
- Computationally difficult to efficiently approximate $p(\theta|m, y, d)$
- Need computationally efficient methods to approximate such probabilities/densities
- Other utilities available e.g: Maximise inverse posterior variance



Sequential designs - SMC single static Model m

Sample from sequence of targets

Data annealing here

$$\pi_t(\theta_m | y_{1:t}, d_{1:t}) = f(y_{1:t} | \theta_m, d_{1:t}) \pi(\theta_m) / Z_{m,t}, \text{ for } t = 1, \dots, T.$$
(1)

 $y_{1:t}$ (independent) data up to t, $d_{1:t}$ design points up to t, θ_m parameter for model m. f is likelihood, π prior and π_t posterior

- SMC: Generate a weighted sample (particles) for each target in the sequence via steps
 - Reweight: particles as data comes in (efficient)
 - Resample: when ESS small
 - Mutation: diversify duplicated particles (can be efficient)



Sequential designs: SMC single static Model *m* (Algorithm) Chopin (2002)

• Have current particles $\{W_t^i, \theta_t^i\}_{i=1}^N$ based on data $y_{1:t}$

Re-weight step to included y_{t+1}

$$W_{t+1}^i \propto W_t^i f(y_{t+1}|\theta_t^i, d_{t+1}),$$

- Check effective sample size: ESS = $1 / \sum_{i=1}^{N} (W_{t+1}^i)^2$
- If ESS > E (e.g. E = N/2) go back to re-weight step for next observation
- If ESS < *E* do the following
- Resample proportional to weights. Duplicates good particles
- Mutation: Move all particles via MCMC kernel say R times (adaptive proposal)

SMC: Estimate of model evidence (Del Moral et al., 2006)

It can be shown

$$Z_{t+1}/Z_t = f(y_{t+1}|y_{1:t}, d_{t+1}) = \int_{\theta} f(y_{t+1}|\theta, d_{t+1}) \pi(\theta|y_{1:t}, D_t) d\theta.$$

 Using SMC particles to approximate posterior at t gives estimator

$$Z_{t+1}/Z_t \approx \sum_{i=1}^N W_t^i f(y_{t+1}|\theta_t^i, d_{t+1}).$$

• Can then obtain approximation of Z_{t+1} through

$$\frac{Z_{t+1}}{Z_0} = \frac{Z_{t+1}}{Z_t} \frac{Z_t}{Z_{t-1}} \cdots \frac{Z_1}{Z_0}$$

Also gives estimate of posterior predictive probability of y_{t+1}

QUI

The Algorithm (Using SMC to design under model uncertainty)

- Effectively run an SMC algorithm for each model m = 1, ..., K
- Have set of *N* particles for each model $\{W_{m,t}^i, \theta_{m,t}^i\}_{i=1}^N$.
- ESS for each model m
- resampling and within-model updates when required
- Design part: use data up to t, $y_{1:t}$, and particles of all models to compute the next design d_{t+1}
- Algorithm has been used by a variety of authors, e.g. discrete choice experiments (Consonni, Deldossi and Saggini, 2018)

QUI

イロト 不得 とくほとく ヨト

Sequential/static designs - Laplace approximation

- Long et al. (2013) and Overstall, McGree and Drovandi (2018) used the Laplace approximation for efficiently estimating u(d, m, y);
- The main result is that the approximation to the posterior distribution of θ_m has the following multivariate Normal form:

$$\hat{\mathsf{p}}(\theta_m|y,d,m) = (2\pi)^{-\frac{q_m}{2}} |\mathbf{\Sigma}_{my}|^{-\frac{1}{2}} \exp\left(-\frac{1}{2}(\theta_m - \hat{\theta}_{my})^t \mathbf{\Sigma}_{my}^{-1}(\theta_m - \hat{\theta}_{my})\right)$$

where q_m denotes the number of parameters in model m, $\hat{\theta}_{my}$ and Σ_{my} denote the posterior mode and posterior variance-covariance matrix, respectively, for model m upon the observation of y from design d, for m = 1, 2, ..., K.

Laplace approximation

- For posterior inference on *m*, the posterior model probability (*p*(*m*|*y*, *d*)) can be considered;
- Proportional to the model evidence;
- Based on Laplace approximation, the model evidence can be approximated as follows:

$$\hat{p}(y|m,d) = (2\pi)^{\frac{q_m}{2}} |\boldsymbol{\Sigma}_{my}|^{\frac{1}{2}} p(y|\hat{\theta}_{my},d) p(\hat{\theta}_{my}|m).$$
(2)

- Thus, posterior summaries such as u(d, m, y) can be evaluated based on the above Laplace approximation facilitating a relatively efficient approximation to u(d);
- Extensions to more complicated models are facilitated by the nested-integrated laplace approximation (Rue, Martino and Chopin, 2009)

Locating Bayesian designs

Difficult optimisation problem

- Noisy and expensive utility function
- Only get realisations from function
- Potentially high dimensional
- Some approaches
 - Exhaustive search
 - Mueller algorithm (Mueller, 1999)
 - Approximate coordinate exchange algorithm (ACE, Overstall and Woods, 2017)

ດມາ

Design and big data

Consider a general regression set up:

- Response data y_i
- *p* predictors (design) *d_i*
- $i = 1, \ldots, N$ observations

Our objective is to avoid the analysis of the big data of size N by selecting a subset of size n

イロト イヨト イヨト イヨト

ດມາ

The algorithm (sequential design)

Our approach:

- 1: Use training sample to obtain $p(\theta)$
- 2: while $n_c \le n$ or analysis aim has not been met (where n_c is the current sample size) **do**
- 3: Find $d_t^* = \arg \max_d u(d)$
- 4: Find x_t in data set such that $||x_t d_t||$ is minimised
- 5: Selected x_t and corresponding y_t , and append to current data set
- 6: Update $p(\theta)$, and remove (x_t, y_t) from data set
- 7: end while

ດມ

イロト 不得 とくほとく ヨト

The algorithm (sequential design)

- Finding an optimal sample of n from N would involve a comparisons across all ^(N) potential designs
- This would be computationally prohibitive
- As an alternative, we first search over *d* ∈ *D*, then search the big data for *x* closely to *d*^{*}
- This is computationally efficient
- However, assumes efficiency is related to distance across all variables

QUI

イロト イポト イヨト イヨト

Potentially sub-optimal but pragmatic

The algorithm (sequential design)

- Line 3 of the algorithm is most challenging
- If covariate space is relatively small, then potentially exhaustive search could be used
- Otherwise, approaches discussed above could be implemented
- This is the most computationally intensive step of the algorithm
- The most limiting factor for general applicability of our algorithm
- This step should not be so computationally intensive that it removes the computational advantages of our approach
- Line 4 is relatively straightforward

Consider mortgage default data

- Data consist of:
 - default: binary variable indicating whether or not the mortgage holder defaulted on the loan
 - creditScore: a credit rating (x1)
 - yearsEmploy: the number of years the mortgage holder has been employed at their current job (x2)
 - ccDebt: the amount of credit card debt (x3)
 - houseAge: the age (in years) of the house (x4)
- For the year 2000, there are 1 million records
- Propose logistic regression model for response variable
- Goal is to determine which covariates are useful for prediction

QUT

- To obtain p(θ), an initial training sample of 5,000 was randomly selected from the data set
- The subsequent posterior distribution was used as the prior for design
- Next we value added to the information gained from the initial learning phase through sequential design
- Here, we implemented the SMC algorithm of Drovandi, McGree and Pettitt (2014)

ດມາ

- An estimation utility was considered $(-\log \det p(\theta|y, d))$
- Importance sampling was used to estimate u(d|y_{1:t-1}, d_{1:t-1})
- To locate the optimal design, an exhaustive search was implemented over all combinations of discrete values of the predictors
- As we are interested in determining which variables are useful for prediction, 95% credible intervals were formed for each slope parameter in every iteration of our algorithm
- If any credible interval was contained within (-tol, tol), then this predictor was dropped from the model
- Values of *tol* were 0.25, 0.5, 0.75, 1.0
- Algorithm iterated until 1,000 observations had been selected from the data set

OD

イロト イヨト イヨト イヨト

Note: n << N</p>

The covariates in the mortgage case study which were deemed useful for prediction based on tol = 0.25, 0.50, 0.75 and 1.00

tol	Remaining covariates		
0.25	x_1, x_2, x_3, x_4		
0.50	x_2, x_3		
0.75	<i>x</i> 3		
1.00	<i>x</i> ₃		

Summary of the posterior distribution of the parameters for the full main effects model based on all mortgage default data for the year 2000

Parameter	Mean	SD	2.5th	Median	97.5th
β_0	-11.40	0.13	-11.67	-11.40	-11.16
β_1	-0.42	0.03	-0.48	-0.42	-0.36
β_2	-0.63	0.03	-0.68	-0.63	-0.56
β_3	3.03	0.05	2.94	3.03	3.13
β_4	0.20	0.03	0.12	0.20	0.26



イロト 不得 とくほとく ヨト



- 212 participants performed a series of 12 different activities at four different time points
- Participants ranged in age between 5 and 18 years
- The purpose was to assess the performance of 4 different so-called "cut-points", which are used to predict the type of activity performed based on the output of the accelerometer
- Response variable was whether or not the cut-point correctly classified the activity
- Each individual at each time point performed all 12 activities and all 4 cut-points are applied

QUI

イロト イヨト イヨト イヨト

Approximately 35,000 data points

A logistic regression mixed effects model was considered
Predictors are age, type of activity (12 levels) and cut-points (4 levels)

ogit
$$\pi_{ti} = \beta_0 + b_t + \beta_1 age_{ti} + \sum_{j=1}^3 \beta_2^j cut_{ti}^j + \sum_{j=1}^{11} \beta_3^j trial_{ti}^j$$

+ $\sum_{j=1}^{33} \beta_4^j cut_{ti}^j \times trial_{ti}^j + \sum_{j=1}^3 \beta_5^j age_{ti}^j \times cut_{ti}^j$
+ $\sum_{j=1}^{11} \beta_6^j age_{ti}^j \times cut_{ti}^j$

where $Y_{ti} \sim \text{Binary}(\pi_{ti})$, $b_t \sim N(0, \phi)$, for t = 1, ..., 212, $i = 1, ..., s_t$ (number of observations on subject *t*), cut_{ti}^j and trial_{ti}^j are dummy variables.

QUT

- Interested in estimating the age effect for correct classification
- Both main and interaction effects, total of 15 parameters
- Ds-optimality considered (similarly defined as in Example 1)
- Initial training sample was found by randomly sampling individuals to obtain (approximately) 500 observations
- Sequential design iterated until 3,000 observations were obtained
- The optimisation (line 3) was a simple grid search over age
- For line 4, we selected the individual who had the closest age to the optimal (48 observations if full replicate was available)

ດມາ

- For fast posterior inference, INLA was used
- Further computational efficiency was achieved by using a small number of Monte Carlo draws
- Only interested in estimating u(d) well enough to determine optimal age

Milan 2018

Thus, can sacrifice precision for computational efficiency

James McGree

ດມາ







QUT

Conclusion

- Proposed an approach to make big data problems more tractable
- Examples showed promise
- Need to consider the trade-off between analysing the big data versus finding the optimal design. Is it worth it?
- Initial work in this space. Many questions/issues remain that we did not tackle

ດມາ

Future research

- Approximate bias in big data sample
- Handle streaming data Design templates?
- Model that can vary/evolve over time?
- Model uncertainty/criticism
- Model-free design?

QUI

Selected references

- Box and Hill (1967). Technometrics, 9, 57-71.
- Chopin (2002). Biometrika, 89, 539-551.
- Del Moral, Doucet and Jasra (2006) JRSS(B), 68, 411-436.
- Drovandi, McGree, and Pettitt (2014) JCGS, 23, 3-24.
- Hill (1978). Technometrics, 20, 15-21.
- Overstall, McGree and Drovandi (2018) Statistics and Computing, 28, 343-358.

ດມາ